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# Evaluating of beef cattle performance and profitability using robust regression analysis

Besi sığırı performansının ve karlılığının dayanıklı regresyon analizi ile değerlendirilmesi

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ABSTRACT

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This work is licensed under a Creative Commons Attribution-Non Commercial 4.0 International License. normal. In this study, we revise Koknaroglu et al. (2005) study in which regression analysis is used under normality assumption. Different from the mentioned study, we use a robust estimation method called M-estimation since the error terms do not follow a normal distribution according to Shapiro-Wilk normality test. We obtain parameter estimates and their standard errors along with the coefficient of determination. It is observed that the results obtained based on M-estimation are more reliable than their LS counterparts with respect to  $R^2$  criterion.

There are several variables affecting the beef cattle performance and profitability. In the

related literature, regression analysis is performed to find the contribution of the variables on

profitability in different managing systems. Least squares (LS) estimation method is mostly

used in regression analysis. However, it is optimal when the distribution of the error terms is

Key Words: Beef cattle, Profitability, Regression, Robustness, M-estimation.

## ÖZ

Besi sığırı performansını ve karlılığın etkileyen birkaç değişken vardır. İlgili literatürde, farklı yönetim sistemlerinde karlılığı etkileyen değişkenlerin katkısını bulmak için regresyon analizi kullanılmıştır. Regresyon analizinde çoğunlukla en küçük kareler (LS) tahmin yöntemi kullanılır. Ancak, bu yöntem hata terimlerinin dağılımının normal olması durumunda optimaldir. Bu çalışmada, normallik varsayımı altında regresyon analizinin kullanıldığı Koknaroğlu ve ark. (2005) çalışması revize edilmiştir. Shapiro-Wilk normallik testine göre hata terimleri için normallik varsayımını sağlanmadığından, diğer çalışmadan farklı olarak, bu çalışmada, M-tahmini adı verilen dayanıklı/robust tahmin yöntemi kullanılmıştır. Belirleme katsayısı ile birlikte parametre tahminleri ve onların standart hataları elde edilmiştir. Sonuç olarak,  $R^2$  kriterine göre, M-tahminine dayalı olarak elde edilen sonuçların, LS tahmin edicilerinden daha güvenilir olduğu gözlenmiştir.

Anahtar Kelimeler: Besi sığırı, Karlılık, Regresyon, Dayanıklılık, M-tahmini.

## Introduction

Regression analysis is widely used statistical technique in most of the applied sciences. The linear regression model is given by

$$y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \ldots + \beta_p x_{ip} + \varepsilon_i, \ i = 1, 2, \dots, n \quad (1)$$

where  $y_i$  is the response,  $x_{i1}, x_{i2} \dots, x_{ik}$  are the predictors,  $\beta_0, \beta_1, \beta_2, \dots, \beta_p$  are the model parameters and  $\varepsilon_i$  denote the random error term. This model can also be written by using alternative representations: (2)

$$y_i = x_i'\beta + \varepsilon_i,$$

and

i = 1, 2, ..., n

where

$$y = [y_1 \ y_2 \ \cdots \ y_n]', \qquad x_i' = \begin{bmatrix} 1 \ x_{i1} \ x_{i2} \ \cdots \ x_{ip} \end{bmatrix}, \qquad \beta = [\beta_0 \ \beta_1 \ \beta_2 \ \cdots \ \beta_k]',$$
$$X = \begin{bmatrix} 1 \ x_1 \ x_2 \ \cdots \ x_p \end{bmatrix}, \qquad 1 = \begin{bmatrix} 1 \ 1 \ \cdots \ 1 \end{bmatrix}' \text{ and } \varepsilon = [\varepsilon_1 \ \varepsilon_2 \ \cdots \ \varepsilon_n]'.$$

The model parameters are estimated using well-known least squares (LS) method. The idea underlying the LS method is to minimize the sum of the squares of error terms with respect to the parameters of interest. Then the well-known LS estimator is formulated by

 $y = X\beta + \varepsilon$ 

$$\hat{\beta} = (X'X)^{-1}X'y. \tag{4}$$

Traditionally, the error terms are assumed to be independently and identically distributed normal with mean zero and variance  $\sigma^2$  in model (1). This assumption has a vital role in statistical inference, i.e. the LS estimators are the most efficient and the test statistics based on them are the most powerful under normality. However, nonnormality is more common in practical studies, see for example Geary (1947), Huber (1981), Islam and Tiku (2005), Senoglu (2005), Koknaroglu et al. (2008), Acitas et al. (2013a, 2013b), Acitas and Senoglu (2019). In the presence of nonnormality, the LS estimators may become inefficient. Furthermore, normal-theory test statistics are not reliable since they are no longer powerful.

Robust statistical methods are frequently used when existing of nonnormality. The aim of the

robust statistical methods is to reduce the large effects of the outliers which cause nonnormality. This is done by giving small weights to the outlying observations. There are several robust estimation methods such as M (Huber; 1964, 1981), MM (Yohai, 1987), least median squares (LMS) (Rousseeuw, 1984), modified maximum likelihood (MML) (Tiku; 1967, 1968) and so on. Mestimation is the most widely used and popular method among them. See also Karadavut and Taşkın (2017) in which M estimation method is used for determination of outliers in Japanese quail body weight data.

In this study, we use model (1) to identify variables that affect beef cattle profitability. Main purpose of animal production is to make living by earning money thus, profitability is the driving force behind animal production. There are several documented variables affecting beef cattle performance and profitability (Koknaroglu et al., 2005). Thus finding contribution of variables on profitability in different managing systems becomes important. Then, the model is formulated by

$$profit_{i} = \beta_{0} + \beta_{1}fedprice + \beta_{2}feedgain + \beta_{3}fe + \beta_{4}corn + \beta_{5}adg + \varepsilon_{i}$$
(5)

where i refers to the pen of cattle. See Koknaroglu et al. (2005) for detailed information on descriptions of the variables used in Model (5).

We revise Koknaroglu et al. (2005) study in which same model is used. Different than the mentioned study, we here use M-estimation method. The reason using M-estimation method is that the error terms do not follow a normal distribution according to Shapiro-Wilk normality test. It should be noted that we do not consider other assumptions regarding to regression model (5) since our aim is to provide a different approach to explore the contribution of several variables on beef cattle profitability in terms of robust regression aspect by revising Koknaroglu et al. (2005).

The rest of the paper is organized as follows. Material and methods are given in section 2. Section 3 is reserved to the results and the discussion. The paper ends with a conclusion section.

#### **Material and Methods**

In this section, data structure, the results of Shapiro Wilk normality test and brief information about the M-estimation method are provided.

## Data structure

Close-out information, consisting of data from cattle that were placed on feed between January 1988 and December 1997, which had been submitted by Iowa cattle producers using the Iowa State University Feedlot Performance and Cost Monitoring program, was examined to determine factors affecting beef cattle performance and profitability.

The model given in equation (5) is employed to identify variables that affect cattle feeding profitability. All of the variables used in model (5) are continuous. Detailed information on materials and methods on how data are obtained and categorized is provided in (Koknaroglu et al., 2005).

The effects of season, housing, sex, body weight (BW) and concentrate level on the profit are also considered. For this purpose, model (5) is applied for different levels of the following variables: season, housing, sex, BW and concentrate level. A brief information about these variables are given as follows. Season has four levels as known well: Winter, spring, summer and fall. Housing includes three levels named confinement, overhead shelter and open lot. Steers and Heifers are two levels of Sex. Initial BW has three levels: cattle weighing <273 kg, between 273 and 364 kg, and >364 kg and finally concentrate level consists of three levels: low (<75%), intermediate (between 75 and 85%), and high (>85%). Therefore, in total 15 regression models should be taken into account during the analysis.

## M-estimation method

M-estimators of the model parameters are solutions of the following minimization problem:

$$\hat{\beta}_M = \underset{\beta}{\operatorname{argmin}} \sum \rho\left(\frac{y_i - x'_i\beta}{\sigma}\right)$$
 (6)

where  $\rho(\cdot)$  is the objective function. After taking derivative in equation (6) with respect to  $\beta$  and setting equal to zero,  $\hat{\beta}_M$  can also be obtained as solution of the following equation:

$$\sum \psi \left( \frac{y_i - x_i' \beta}{\sigma} \right) x_i = 0 \tag{7}$$

where  $\psi(\cdot) = \rho'(\cdot)$ . It should be noted that LS estimators are obtained when  $\rho(x) = x^2$ . Different choices of  $\rho(\cdot)$  function are considered to capture the robustness in the related literature. Indeed,  $\rho(\cdot)$  function should satisfy some properties for the sake of robustness. These properties are given in Maronna et al. (2006) on page 31. Jureckova and Picek (2010) can also be seen for different  $\rho(\cdot)$  functions.

In this study, we use Tukey's bisquare (biweight) function:

$$\rho(x) = \begin{cases} 1 - \left[1 - \left(\frac{x}{k}\right)^2\right]^3 & if \quad |x| \le k \\ 1 & if \quad |x| > k. \end{cases}$$
(8)

Therefore,  $\psi(\cdot)$  function is obtained by

$$\psi(x) = \begin{cases} \frac{6}{k^2} x \left[ 1 - \left(\frac{x}{k}\right)^2 \right]^2 & if \quad |x| \le k \\ 0 & if \quad |x| > k \end{cases}$$
(9)

Here, k is the robustness tuning constant which is used for adjusting the trade-off between the robustness and the efficiency. Therefore, it is taken as 4.68. Tukey's bisquare function is mostly used in robust statistical analyzes since  $\psi$  function is redescending, i.e.  $\psi$  tends to 0 for  $\infty$ . The advantage of using a redescending  $\psi$  function is that it provides more robustness (Maronna et al., 2006). The plots of  $\rho(x)$  and  $\psi(x)$  functions are given in Figure 1.

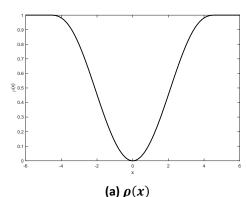


Figure 1. Plots of  $\rho(x)$  and  $\psi(x)$  functions Şekil 1.  $\rho(x)$  ve  $\psi(x)$  fonksiyonlarının grafikleri

It is clear that solutions of equation (7) cannot be obtained explicitly. Therefore, numerical methods should be performed. In the related literature, iteratively reweighted least squares (IRLS) method is mostly utilized to compute the M-estimates, see for example Montgomery (2012). The computations in this study are done using "robustfit" function of MATLAB software. Robustfit function gives M-estimates of the model parameters in addition to robust estimate of the scale. We also compute the robust coefficient of determination ( $R^2$ ) given by

$$R_w^2 = 1 - \frac{\sum w_i (y_i - \hat{y}_i)^2}{\sum w_i (y_i - \bar{y}_w)^2}$$
(10)

where

$$ar{y}_w = rac{1}{\sum w_i} \sum w_i y_i, \quad \hat{y}_i = y_i - x'_i \, \hat{eta}_M$$

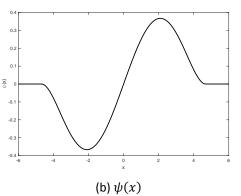
and  $w_i$  (i = 1, 2, ..., n) are the weights, see Renaud and Victoria-Fesser (2010). Obviously, the higher values of  $R^2$  implies better fitting.

## **Results and Discussion**

In this part of the study, we first explore the normality of the data set. Then we analyze the data set using the M-estimation method and interpret the results.

## Normality test results

In this study, we use the Shapiro-Wilk (1965), one of the widely used and the most powerful goodness of fit test, to explore the normality of the data set. The appropriate hypotheses are



stated as follows:

Ho: The error terms have a normal distribution H1: The error terms do not have a normal distribution

The test is conducted using LS residuals. Indeed, we first compute the LS estimates of the model parameters using equation (4) and then obtain the residuals based on them. The test is conducted in MATLAB using Oner and Deveci's (2017) code file which is available at the website provided in references.

As it is indicated previously, there are five variables. Model (5) is used for each level of these variables. Therefore, we have 15 regression models belonging to the different levels of season, housing, sex, body weight and concentrate level. Shapiro-Wilk test is carried out to LS residuals obtained from these 15 regression models. In other words, first LS residuals are obtained for each 15 regression models. Then, they are used to check the normality assumption via Shapiro-Wilk test. The results are given in Table 1. It reports value of the test statistics and the corresponding p –values.

The results are interpreted at  $\alpha = 0.05$  significance level as follows. It is clear that the normality assumption is not satisfied for Winter, Spring and Summer while it is satisfied for Fall. For three levels of housing variable, the normality of the error terms is rejected. While the error terms are distributed normally for Heifers, the distribution is not normal for Steers. First two levels of BW, the normality assumption is satisfied. However, third level (>364) does not

follow a normal distribution. For the concentrate levels (75 to 85 and >85), the normality

assumption is met. On the other hand, for the level <75, the null hypothesis is rejected.

#### Table 1. Shapiro-Wilk normality test results *Cizelge 1. Shapiro-Wilk normallik testi sonucları*

		Sample Size (Örneklem hacmi)	Test Statistic (Test İstatistiği)	p —value ( $p$ —değeri)	Not reject/Reject (Red değil/Red) (0/1)	
Season (Mevsim)	Winter (Kış)	448	0.9798	0.0001	1	
	Spring (İlkbahar)	300	0.9559	0.0001	1	
	Summer (Yaz)	420	0.9769	0.0001	1	
	Fall (Sonbahar)	685	0.9964	0.1171	0	
Housing (Barınma)	Confinement (Kapalı ahır)	456	0.9089	0.0001	1	
	Overhead shelter (Yarı açık ahır)	470	0.9930	0.0273	1	
	Open lot (Açık ahır)	927	0.9929	0.0002	1	
Sex (Cinsiyet)	Steers (Erkek Dana)	1429	0.9834	0.0001	1	
	Heifers (Dişi Dana)	424	0.9942	0.1046	0	
BW (Vücut Ağırlığı)	<273	371	0.9832	0.0003	1	
	273 to 364	964	0.9699	0.0001	1	
	>364	518	0.9947	0.0713	0	
Concentrate Level (Yoğunluk Düzeyi)	<75	180	0.9895	0.2062	0	
	75 to 85	1020	0.9928	0.0001	1	
	>85	653	0.9505	0.0001	1	

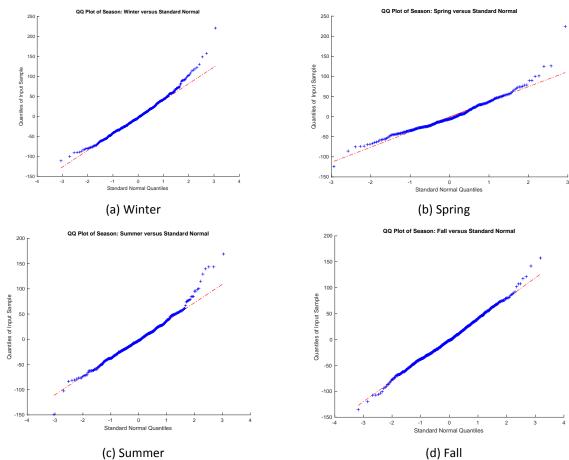


Figure 2. Q-Q plots for season Şekil 2. Mevsim için Q-Q grafikleri

These results are also supported by Figures 2 - 6 in which Q-Q plots are given for all categories. It should be noted that the normality assumption is

satisfied if the quantile pairs do not deviate too much from the straight line.

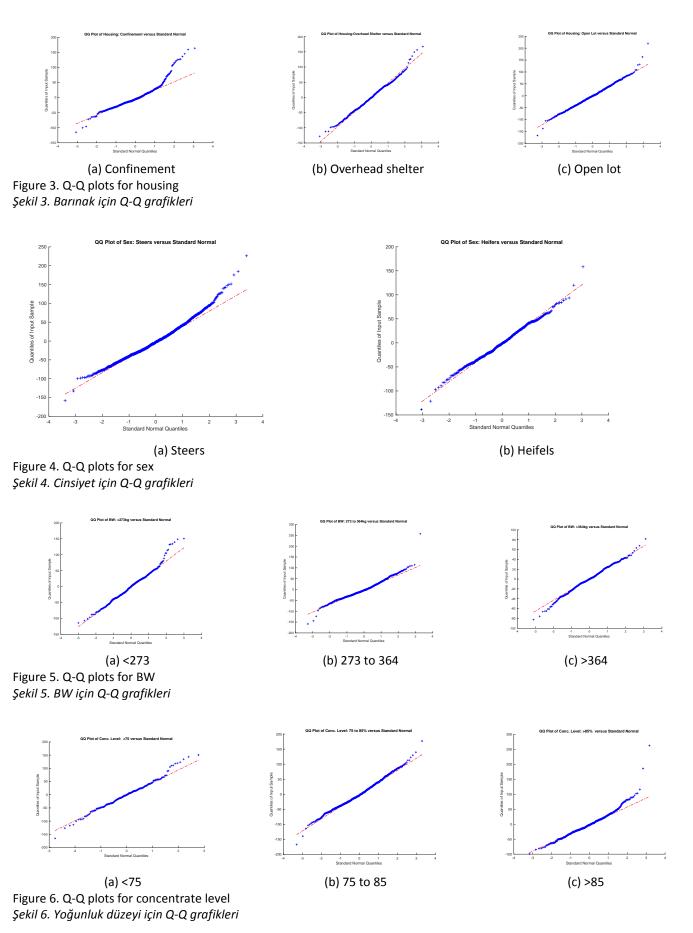


Table 2. Estimated regression coefficients of the factors explaining profitability,  $\hat{\sigma}$  and  $R^2$ . The values given in parenthesis are the standard error of the regression estimates

Çizelge 2. Karlılığı açıklayan faktörler için tahmin edilmiş regresyon katsayıları,  $\hat{\sigma}$  ve  $R^2$ . Parantez içinde verilen değerler regresyon tahminlerinin standart hatalarıdır

				$\hat{eta}_0$	$\hat{eta}_1$	$\hat{eta}_2$	$\hat{eta}_3$	$\hat{eta}_4$	$\hat{eta}_5$	$\hat{\sigma}$	R <sup>2</sup> (%)
		LS	Coeff.	90.28	9.76	-4.90	-27.98	-29.65	-64.57	45.49	61
	Winter		SE	(63.46)	(0.43)	(0.29)	(3.36)	(5.55)	(17.02)		
	(Kış)	М	Coeff.	54.51	9.98	-5.12	-25.96	-27.52	-52.49	44.43	68
			SE	(62.10)	(0.42)	(0.28)	(3.28)	(5.43)	(16.65)		
		LS	Coeff.	86.54	10.46	-6.01	-27.90	-34.63	-25.22	39.63	67
	Spring (İlkbahar)		SE	(66.36)	(0.54)	(0.32)	(3.17)	(6.66)	(17.03)		
		М	Coeff.	60.65	11.11	-6.76	-24.74	-45.38	4.12	36.31	77
Season			SE	(61.41)	(0.50)	(0.30)	(2.93)	(6.17)	(15.76)		
(Mevsim)		LS	Coeff.	-99.89	11.50	-5.40	-24.73	-21.85	-26.55	40.36	71
	Summer		SE	(56.74)	(0.43)	(0.23)	(2.45)	(6.66)	(11.91)		
	(Yaz)	М	Coeff.	-93.99	12.37	-6.76	-20.67	-34.25	0.52	34.32	84
			SE	(48.98)	(0.37)	(0.20)	(2.12)	(5.75)	(10.28)		
		LS	Coeff.	230.99	8.60	-4.68	-34.61	-41.47	-58.06	40.52	67
	Fall		SE	(40.24)	(0.30)	(0.19)	(1.85)	(5.19)	(10.08)		
	(Sonbahar)	М	Coeff.	235.02	8.64	-4.75	-35.19	-40.94	-57.56	40.72	72
			SE	(40.44)	(0.30)	(0.19)	(1.86)	(5.21)	(10.13)		
		LS	Coeff.	207.38	9.60	-6.15	-31.94	-41.75	-34.17	37.02	74
	Confinement		SE	(49.84)	(0.34)	(0.21)	(2.84)	(6.59)	(12.51)		
	(Kapalı ahır)	М	Coeff.	136.07	10.29	-6.95	-27.72	-45.67	9.30	26.24	89
			SE	(36.59)	(0.25)	(0.15)	(2.09)	(4.84)	(9.18)		
	Overhead shelter (Yarı açık	LS	Coeff.	74.29	9.99	-4.45	-34.69	-20.65	-68.17	45.63	61
Housing			SE	(62.99)	(0.42)	(0.28)	(2.88)	(6.52)	(15.30)		
(Barınma)		М	Coeff.	47.67	10.19	-4.53	-33.96	-20.29	-58.43	45.68	66
	ahır)		SE	(63.06)	(0.42)	(0.28)	(2.89)	(6.53)	(15.32)		
		LS	Coeff.	117.66	9.44	-4.98	-30.09	-30.55	-46.86	41.44	66
	Open lot		SE	(37.15)	(0.27)	(0.18)	(1.63)	(3.71)	(8.72)		
	(Açık ahır)	М	Coeff.	122.75	9.54	-5.19	-29.88	-31.21	-43.74	41.41	71
			SE	(37.12)	(0.27)	(0.18)	(1.63)	(3.71)	(8.72)		
Sex (Cinsiyet)		LS	Coeff.	165.34	9.76	-5.22	-34.21	-33.06	-60.76	43.34	65
	Steers		SE	(33.23)	(0.23)	(0.15)	(1.60)	(3.48)	(8.17)		
	(Erkek Dana)	М	Coeff.	155.72	10.02	-5.52	-34.00	-33.22	-51.59	42.23	72
			SE	(32.39)	(0.22)	(0.15)	(1.56)	(3.39)	(7.97)		
		LS	Coeff.	3.38	9.33	-4.41	-25.26	-20.04	-36.36	38.77	66
	Heifers		SE	(49.28)	(0.39)	(0.24)	(2.03)	(5.33)	(14.53)		
	(Dişi Dana)	N.4	Coeff.	-20.61	9.42	-4.41	-24.23	-19.32	-30.24	38.96	70
	(Dişi Dana)	Μ	coen.	-20.01	9.42	-4.41	-24.23	-19.52	-30.24	50.90	70

## Table 2. (Cont.) *Çizelge 2. (Devamı)*

				$\hat{eta}_0$	$\hat{eta}_1$	$\hat{eta}_2$	$\hat{eta}_3$	$\hat{eta}_4$	$\hat{eta}_5$	ô	R <sup>2</sup> (%)
	<273 —	LS	Coeff.	-132.46	9.94	-3.84	-21.84	-37.18	35.73	44.08	64
			SE	(64.76)	(0.46)	(0.22)	(3.20)	(6.79)	(21.66)		
		М	Coeff.	-124.52	10.10	-3.85	-23.05	-42.95	38.39	42.83	70
			SE	(63.09)	(0.45)	(0.22)	(3.12)	(6.61)	(21.10)		
		LS	Coeff.	42.37	10.30	-5.86	-25.79	-39.63	5.93	36.48	74
BW	273 to 364 —		SE	(33.54)	(0.23)	(0.16)	(1.68)	(3.40)	(9.43)		
(Vücut Ağırlığı)		М	Coeff.	22.60	10.52	-6.04	-24.20	-40.59	11.51	35.10	80
			SE	(32.32)	(0.22)	(0.16)	(1.62)	(3.28)	(9.09)		
	>364 –	LS	Coeff.	-16.96	11.42	-8.03	-16.97	-35.58	39.97	23.72	89
			SE	(31.54)	(0.22)	(0.16)	(1.42)	(3.43)	(7.77)		
		М	Coeff.	-0.99	11.57	-8.32	-16.55	-40.57	43.56	23.62	91
			SE	(31.41)	(0.22)	(0.16)	(1.42)	(3.42)	(7.74)		
	<75 –	LS	Coeff.	-19.63	10.86	-4.57	-29.20	-29.41	-33.59	53.36	60
			SE	(94.89)	(0.81)	(0.40)	(4.05)	(10.64)	(23.47)		
Concentrate Level (Yoğunluk Düzeyi)		М	Coeff.	-100.44	11.40	-4.52	-27.25	-27.50	-19.91	51.86	67
			SE	(92.68)	(0.79)	(0.39)	(3.95)	(10.39)	(22.92)		
	75 to 85 —	LS	Coeff.	104.35	9.70	-4.96	-31.08	-29.23	-48.68	40.59	67
			SE	(35.56)	(0.25)	(0.16)	(1.69)	(3.74)	(9.03)		
		М	Coeff.	83.55	9.69	-4.93	-30.18	-28.47	-41.02	40.73	71
			SE	(35.68)	(0.25)	(0.16)	(1.70)	(3.75)	(9.07)		
	>85 -	LS	Coeff.	159.23	9.88	-5.94	-32.44	-39.70	-26.98	35.62	74
			SE	(38.59)	(0.27)	(0.19)	(1.93)	(4.35)	(9.55)		
		М	Coeff.	172.36	10.43	-7.18	-28.60	-48.36	-1.34	29.18	87
			SE	(32.02)	(0.23)	(0.16)	(1.60)	(3.61)	(7.92)		

## Regression analysis results

This section includes LS and M-estimates of the model parameters. Estimated regression coefficients of the variables explaining profitability are given in Table 2. Furthermore, LS and robust estimates of scale parameter (  $\hat{\sigma}$  ) and the coefficient of determination  $(R^2)$  are provided in Table 2. The standard errors (SEs) are also given under the corresponding regression estimate in parenthesis. It should also be noted that robust estimate of scale parameter and standard errors for M-estimates of regression parameters are obtained from robustfit function. Following results can be deduced from Table 2.

## Season

The normality assumption is not satisfied for Winter, Spring and Summer. Therefore, standard errors of the M-estimates are less than their LS counterparts. Furthermore,  $\hat{\sigma}_M < \hat{\sigma}_{LS}$  implies that robust linear regression model is more reliable. This conclusion is also supported by  $R^2$ values, i.e.  $R_M^2 > R_{LS}^2$ . For the Fall, normality assumption is valid. Therefore, LS results are preferable here.

## Housing

The normality assumption violated for three levels of housing variable. Therefore, M-estimates are more reliable for this case. This conclusion is obtained from the standard errors of the estimated regression coefficients and the standard deviation of the error terms.  $R_M^2 > R_{LS}^2$  also implies to use robust statistical methods.

## Sex

The distribution of the error terms is not normal for Steers. Therefore, robust statistical

methods should be used. The results show that standard errors of the M-estimates are less than those of LS estimates. However, LS estimates are more preferable for Heifers since normality assumption is satisfied for this case.

## BW

First two levels of BW do not satisfy the normality assumption. Therefore, M-estimation method should be used to estimate the model parameters for these levels. Indeed, standard errors for M-estimates are smaller than their LS counterparts. This is also true for estimate of the scale parameter. The coefficient of determination for M-estimation is also higher. However, for the third level, LS estimates are more preferable since normality is satisfied.

## Concentrate level

The normality assumption is satisfied for the first level (<75). Therefore, LS estimates should be preferred for this case. On the other hand, for the remaining levels (75 to 85% and >85%), the normality is not satisfied. Indeed, the M-estimates are more reliable for these cases.

## Conclusion

Regression analysis is a widely used method in animal sciences. For example, the beef cattle performance and profitability can be determined using the regression analysis. The parameters of the linear regression model are frequently estimated using LS method. It is a well-known fact that LS method is optimal when the distribution of the error terms is normal. However, nonnormality of the error terms is more common in practice. The motivation for this study comes from this fact. Koknaroglu et al. (2005) evaluate the beef cattle performance using linear regression model under normality. However, in our analyzes, we find that normality is not satisfied for most of the cases. We therefore use robust regression methods, i.e. M-estimation. The advantage of using a robust method is that it is not sensitive to the outliers and also to the

nonnormality. Therefore, it gives more reliable results in presence of outliers and nonnormality.

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